

Adaptive Estimation of Partially Identified Treatment Effects

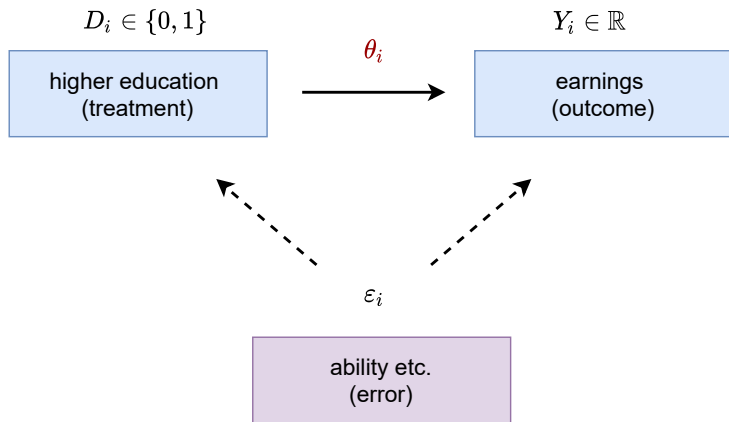
Maria Nareklishvili

The University of Oslo
maria.nareklishvili@frisch.uio.no

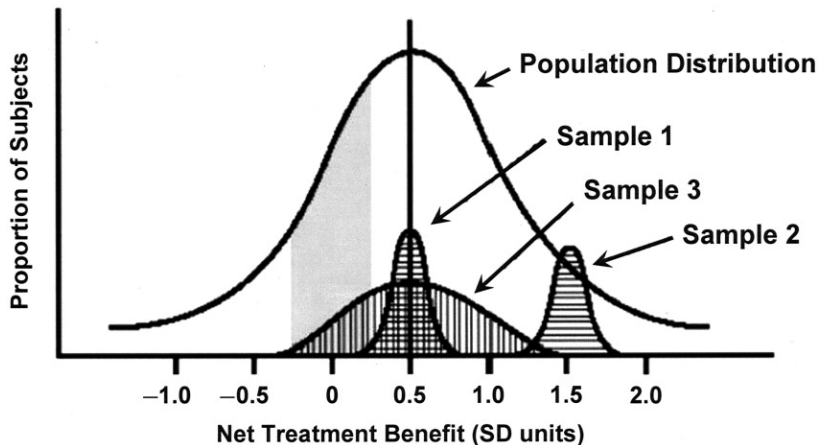
December 11, 2022

Motivation: Endogeneity

Endogenous Treatment



Heterogeneous Treatment Effects



Research Questions

Impose **causal assumptions** to deal with endogeneity and use **random forests** to infer partially identified treatment effects.

- *Conditional treatment effect bounds* tighter and more informative?
- Do conditional bounds reflect *heterogeneous subgroups*?
- Statistically and economically *significant* heterogeneity?

Contribution

- Introduce **flexible direction** of the monotonicity assumptions to identify conditional treatment effect bounds [Manski, 1990; Balke and Pearl, 1997; Heckman and Vytlacil, 1999; Horowitz and Manski, 2000; Beresteanu and Manski, 2000; Chernozhukov et al., 2007; Manski and Pepper, 2009; Lee, 2009; Romano and Shaikh, 2010; Shaikh and Vytlacil, 2011; Beresteanu et al., 2011; Huber and Mellace, 2015; Mogstad and Torgovitsky, 2018; Torgovitsky, 2019; Semenova, 2020; Heiler, 2022].
- Introduce **multivariate random forests** and investigate large sample properties [Breiman, 2001, 2004; Lin and Jeon, 2006; Meinshausen and Ridgeway, 2006; Biau, 2012; Denil et al., 2014; Wager, 2014; Scornet et al., 2015; Athey and Imbens, 2016; Wager and Athey, 2018; Athey et al., 2019; Nekipelov et al., 2018; Chernozhukov et al., 2017, 2018; Li, 2020].

Illustrative Example

National Longitudinal Survey of Youth, 1979 (De Haan and Leuven, 2020). Treatment: the preschool program; outcome: years of schooling.

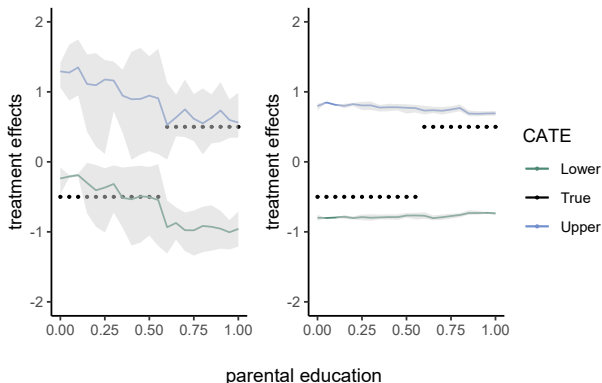


Figure 1: **negative** treatment selection, subgroups recovered by **causal forests** (left), **flexible** treatment selection, subgroups recovered by **MRF** (right).

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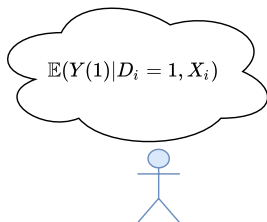
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Identification: Monotonic Treatment Selection

1. Bounded support of the outcome

2. MTS

Observed



A treatment is a preschool program, designed for children with a disadvantaged family background

Unobserved

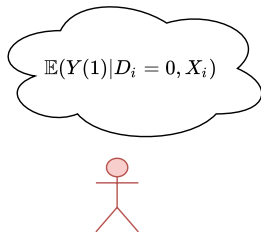
positive selection

\geq

or

negative selection

\leq



Identification: Bounds of CATE

Proposition 1.1

Under a negative treatment selection (Manski, 1990; Manski and Pepper, 2000):

$$\begin{aligned}\theta^L(x) &= \mathbb{E}(Y_i(1)|D_i = 1, X_i = x) - \mathbb{E}(Y_i(0)|D_i = 0, X_i = x) \\ \theta^U(x) &= \mathbb{E}(Y_i(1)|D_i = 1, X_i = x) \cdot P(D_i = 1|X_i = x) + \\ &\quad Y^U(x) \cdot P(D_i = 0, X_i = x) - \mathbb{E}(Y_i(0)|D_i = 0, X_i = x) \cdot \\ &\quad \cdot P(D_i = 0|X_i = x) - Y^L(x) \cdot P(D_i = 1|X_i = x).\end{aligned}$$

Proof

$Y^B(x)$ — user-specified quantiles for $B \in \{L, U\}$.

Assumptions

Assumption 2.1 (Honesty)

$\mathcal{F}(Y_{im}|X_i, s) = \mathcal{F}(Y_{im}|X_i)$. $s = (j, c)$ denotes the splitting coordinates and the splitting indices.

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The probability that the next split occurs at j -th covariate is bounded below by π/p for $\pi \in (0, 1]$, for all $j = 1, \dots, p$.

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Assumption 2.2 (Random Split Trees)

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Assumption 2.3 (The Splitting Algorithm is (α, k) -regular)

Each leaf contains at least $\alpha > 0$ fraction of observations. Moreover, the splitting ceases when $N_\ell \leq k$ for some $k \in \mathbb{N}$.

Assumptions

Assumption 2.4 (Distributional Assumptions on the Data Generating Process)

$X_i \in [0, 1]^p$, $\mathbb{E}(Y_{im}|X_i = x)$, $\mathbb{E}((Y_{im})^2|X_i = x)$ are *Lipschitz-continuous*. Furthermore, $\text{Var}(Y_{im}|X_i = x)$ is bounded away from 0 (i.e., $\inf_{x \in \mathcal{X}} \text{Var}(Y_{im}|X_i = x) > 0$).

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Assumption 2.5 (Overlap)

We assume that for all $x \in [0, 1]^p$:

$$0 < \mathbb{P}(D_i = 1|X_i = x) < 1.$$

Inference: Hoeffding Decomposition I

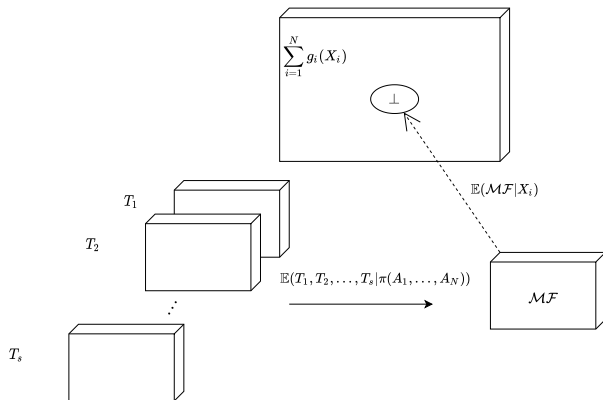


Figure 2: Hajek Projection (Hájek, 1968; Athey and Wager, 2019) of a (vector-valued) U-statistic.

Inference: Hoeffding Decomposition II

More formally,

$$\underbrace{\Sigma^{-1/2}(\mathcal{M}\mathcal{F} - \mu)}_Y = \underbrace{\Sigma^{-1/2}(\dot{\mathcal{M}}\mathcal{F} - \mu)}_{\mu(X_i)} + \underbrace{\Sigma^{-1/2}(\mathcal{M}\mathcal{F} - \dot{\mathcal{M}}\mathcal{F})}_{\varepsilon},$$

The objective:

$$\Sigma^{-1/2}(\mathcal{M}\mathcal{F}(x, A_1, \dots, A_N) - \dot{\mathcal{M}}\mathcal{F}(x, A_1, \dots, A_N)) \xrightarrow{p} 0.$$

Lemma 1

The mean squared difference of $\mathcal{M}\mathcal{F}$ and $\dot{\mathcal{M}}\mathcal{F}$ has the upper bound:
$$\mathbb{E}(\mathcal{M}\mathcal{F} - \dot{\mathcal{M}}\mathcal{F})^T \Sigma^{-1} (\mathcal{M}\mathcal{F} - \dot{\mathcal{M}}\mathcal{F}) \leq \frac{s}{N} \text{tr} \left((\mathbb{V}(\hat{\tilde{T}}))^{-1} \mathbb{V}(\tilde{T}) \right),$$
 where $\mathbb{V}(\hat{\tilde{T}})$ and $\mathbb{V}(\tilde{T})$ denote the variance of the tree projection and a tree, respectively. Sketch of Proof

Inference: Asymptotic Theory

Lemma 1

The mean squared difference of \mathcal{MF} and $\dot{\mathcal{M}}\mathcal{F}$ has the upper bound:
$$\mathbb{E}(\mathcal{MF} - \dot{\mathcal{M}}\mathcal{F})^T \Sigma^{-1} (\mathcal{MF} - \dot{\mathcal{M}}\mathcal{F}) \leq \frac{s}{N} \text{tr} \left((\mathbb{V}(\dot{\tilde{T}}))^{-1} \mathbb{V}(\tilde{T}) \right),$$
 where $\mathbb{V}(\dot{\tilde{T}})$ and $\mathbb{V}(\tilde{T})$ denote the variance of the tree projection and a tree, respectively. [Sketch of Proof](#)

Theorem 2

The entries of $\mathbb{V}(\tilde{T})$ are bounded and its diagonal elements are bounded away from zero. Moreover, the lower bound of the off-diagonal terms of $\mathbb{V}(\tilde{T})$ are on the order of $o\left(\frac{1}{\log^p(s)}\right)$. The upper bound in Lemma 1 converges to zero in the limit: $\frac{s}{N} \text{tr} \left((\mathbb{V}(\dot{\tilde{T}}))^{-1} \mathbb{V}(\tilde{T}) \right) \rightarrow 0$. [Sketch of Proof](#)

Estimation: Method of Moments

Proposition 2.1

Deviation between the personalized and group level parameters:

$$\mathbb{E} \left[(\theta(X_i) - \tilde{\theta}(X_i, S^{est}, \Pi))^T \Sigma^{-1} (\theta(X_i) - \tilde{\theta}(X_i, S^{est}, \Pi)) \right].$$

Quantile loss:

$$\sum_b \mathbb{E}(L_{\alpha_b}(Y_i, q) | X_i).$$

Sketch of Proof

where

$$L_{\alpha_b}(y, q) = \begin{cases} \alpha_b \times |y - q| & \text{if } y > q \\ (1 - \alpha_b) \times |y - q| & \text{if } y \leq q \end{cases}.$$

Simulation: NLSY 1979

Y_i is the years of schooling. D_i - Head Start program.

$$Y_i = 13.240 - 0.5 \cdot D_i + \\ 1\{\text{Parental Education}_i \geq 0.6\} \cdot D_i + X_i^R \cdot 0 + \varepsilon_i, \\ \varepsilon_i \sim \mathcal{N}(0, 1 + D_i).$$

Simulate 100 experiments with the upper quantiles of the outcome within $[1, 0.7]$, and the lower quantiles within $[0, 0.3]$.

Simulations: A Tree

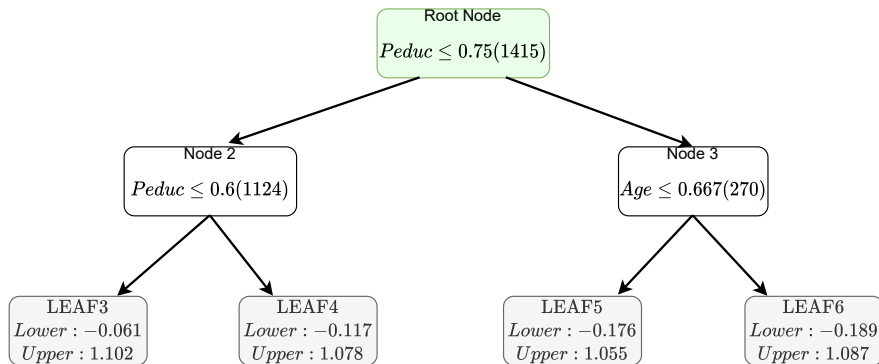
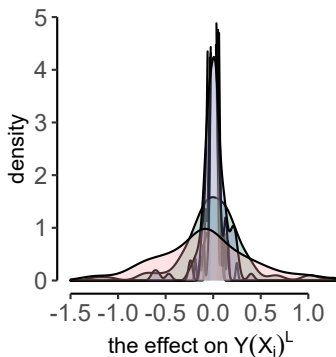
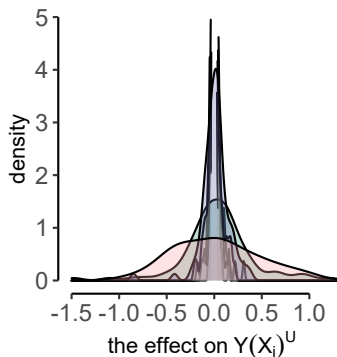


Figure 3: A single (randomly chosen) tree out of 200 trees.

Simulation: Splitting Accuracy

The correlation of parental education with age, the number of siblings, black, Hispanic and female have is -39%, -43%, -5%, 3%, and 2%, respectively.



Application: Treatment Effect Bounds

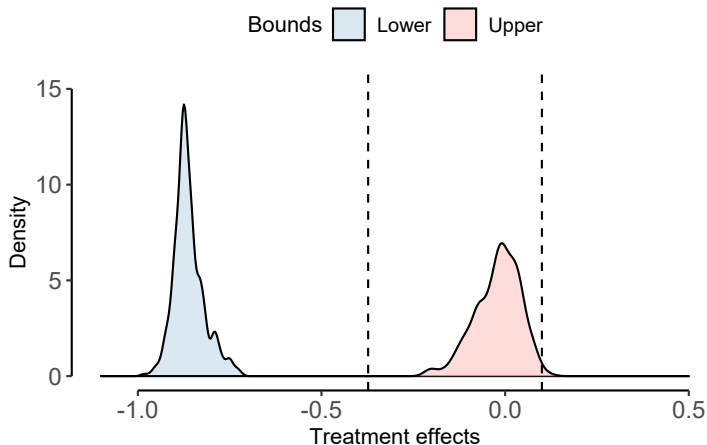
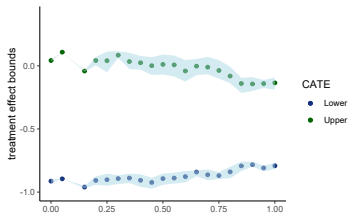
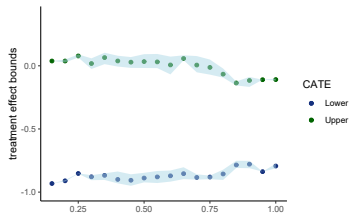


Figure 4: The density of the bounds for the effect of Head Start participation on years of schooling. We use 60 and 30-th percentiles of the outcome. Dashed lines represent unconditional bounds.

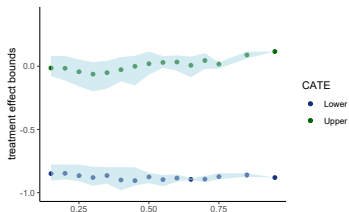
Application: Heterogeneity



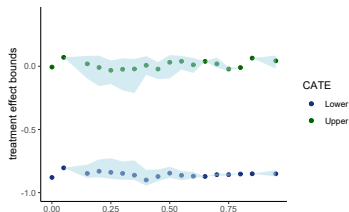
(a) P.education for female



(b) P.education for male



(c) N.Sibs for female



(d) N.Sibs for male

Future Work

This paper introduced

- Multivariate forests for partially identified CATE with an endogenous treatment,
- A test for heterogeneity - [See](#)
- Extension to anomalous data,
- Extension to network data.

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Empirical extensions:

- Investigate personalized effects of the education on earnings induced by the [compulsory schooling reform](#) in Norway (1959-1974).
- Investigate personalized effects of various [cancer](#) treatment methods in Norway.

Thank You!

E-mail: maria.nareklivili@frisch.uio.no

Twitter: <https://twitter.com/mnareklivili>